# Supplementary Material for: <br> Do We Exploit all Information for Counterfactual Analysis? Benefits of Factor Models and Idiosyncratic Correction 

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October 2, 2021


#### Abstract

In this Supplementary Material we provide a number of additional results to the paper "Do We Exploit all Information for Counterfactual Analysis? Benefits of Factor Models and Idiosyncratic Correction". In addition to the proof of the theoretical results in the above mentioned paper, we also report additional empirical results. JEL Codes: C22, C23, C32, C33. Keywords: counterfactual estimation, synthetic controls, ArCo, treatment effects, factor models, high-dimensional testing, LASSO, optimal pricing, retail, price setting, demand. Acknowledgments: We wish to thank an associate editor and three anonymous referees for very insightful comments. Fan's research was supported by NSF grants DMS-1712591, DMS-2052926, DMS-2053832 and ONR grant N00014-19-1-2120. Masini's and Medeiros' research was partially supported by CNPq and CAPES. We are also in debt with Thiago Milagres for helping us with the dataset and all the team from the D-LAB@PUC-Rio for providing a superb research environment.


## 1 Introduction

This is a supplement to the paper "Do We Exploit all Information for Counterfactual Analysis? Benefits of Factor Models and Idiosyncratic Correction". The document is organized as follows. In Section 2 we describe the algorithm used to split the cities into the treatment and control groups. Section 3 contains additional empirical results. More specifically, in Section 3.2 we compare the empirical results when the ArCo methodology of Carvalho, Masini, and Medeiros (2018) and the Principal Component Regression (PCR) as in Gobillon and Magnac (2016) are used to estimate the counterfactuals. In Section 3.3 we evaluate different approaches to model trending behavior in the data, while in Section 3.4 we present the results at a state-level aggregation. The proof of the main result in the paper is presented in Section 4

## 2 Randomization Algorithm

In this section we describe the algorithm used to split the municipalities into two different groups according to a set of characteristics. Once the groups are formed we randomly label them as treatment and control groups.

Let $\boldsymbol{Z}$ be a $n \times J$ matrix of municipalities' variables, where each column $j$ is a different characteristic (covariate) and each row $i$ is a municipality, $n$ is the number of municipalities and $J$ is the number of covariates. We consider the following variables: human development index, employment, GDP per capita, population, female population, literate population, average household income (total), household income (urban areas), number of stores, and number of convenience stores.

The goal is to match the average of each characteristic of the treatment group with the control group. Once each group of municipalities is created, each group is further divided into two other groups, resulting in four different sets of municipalities. The experiments were carried on different combinations of the groups. In the paper, we report only one set of the experiments.

The optimization problem is defined as

$$
\begin{aligned}
& \hat{\boldsymbol{\alpha}}=\arg \min _{\boldsymbol{\alpha}} \frac{1}{J} \sum_{j=1}^{J}\left|\frac{1}{\sum_{i=1}^{n} \alpha_{i}} \sum_{i=1}^{n} \alpha_{i} Z_{i, j}-\frac{1}{\sum_{i=1}^{n}\left(1-\alpha_{i}\right)} \sum_{i=1}^{n}\left(1-\alpha_{i}\right) Z_{i, j}\right| \\
& \text { subject to: } \sum_{i=1}^{n} \alpha_{i}=K \text { and } \alpha_{i} \in\{0,1\} \forall i,
\end{aligned}
$$

where $\boldsymbol{\alpha}=\left(\alpha_{1}, \ldots, \alpha_{n}\right)^{\prime}, \alpha_{i}=1$ if municipality $i$ belongs to the first group and $\alpha_{i}=0$ otherwise; $K$ is the number of municipalities in the first group. The optimization problem above can be transformed into a mixed integer program.

## 3 Additional Empirical Results

In this section we report a number of additional empirical results with the aim of showing the robustness and advantages of the proposed methodology.

### 3.1 Additional Plots

Figures S. 1 S. 4 display relevant data for Products II-V. Panel (a) in the figures reports the daily sales at each group of municipalities (all, treatment, and control) divided by the number of stores in each group. More specifically, the plot shows the daily evolution of $q_{\text {all, }, t}^{(j)}=\frac{1}{s} \sum_{i=1}^{n} q_{i t}^{(j)}$, $q_{\text {control }, t .}^{(j)}=\frac{1}{s_{0}} \sum_{i=1}^{n_{0}} q_{i t}^{(j)}$, and $q_{\text {treatment }, t .}^{(j)}=\frac{1}{s_{1}} \sum_{i=n_{0}+1}^{n} q_{i t}^{(j)}$. The plot shows the data before and after price changes and the intervention date is represented by the horizontal line. Panels (b) and (c) display the distribution across municipalities of the time averages of $\widetilde{q}_{i t}^{(j)}$, before and after the intervention and for the treatment and control groups, respectively. Panels (d) and (e) present fan plots for the evolution of $\widetilde{q}_{i t}^{(j)}$. The black curves there represent the cross-sectional means over time.

Figures S.5 S. 8 display some estimation results. For each product, Panel (a) in the figures displays a fan plot of the $p$-values of the ressampling test for the null hypothesis $\mathcal{H}_{0}: \delta_{t}=0$ for each given $t$ after the treatment, using the test statistic $\phi\left(\hat{\delta}_{t}\right)=\left|\hat{\delta}_{t}\right|$, which is the same as using the test statistic $\hat{\delta}_{t}^{2}$. The black curve represents the cross-sectional median across time $t$. Panel (b) shows an example for one municipality. The panel shows the actual and counterfactual sales per store for the post-treatment period. $95 \%$ confidence intervals for the counterfactual path are also displayed.

Figure S.9 displays the distribution of the daily evolution of the inventory of each product across different municipalities.

### 3.2 Effects of Additional Information: ArCo, PCR, and FarmTreat

We report the estimation results when either the ArCo methodology of Carvalho, Masini, and Medeiros (2018) or principal component regression (PCR) in the spirit of Gobillon and Magnac (2016) are used. For the ArCo methodology we construct counterfactuals by estimating a LASSO regression of $\widetilde{q}_{i t}^{(j)}$ on the values of $\widetilde{q}_{k t}^{(j)}$, where $k \in\{1, \ldots, n\} / i$. Note that we do not include any other regressor. For PCR, we consider the first two stages of the FarmTreat methodology.

The ArCo results are reported in Tables S.1 and S.2 while the results for the PCR method are shown in Tables S.3 and S.4. Some interesting facts emerge from the tables. First, the ArCo and FarmTreat show similar results, with the later having a slightly better pre-intervention fit. One key difference, however, is the substantially smaller number of municipalities with significant intervention effects when the ArCo methodology is considered. Comparing the PCR approach and the FarmTreat, we can clearly see an improvement in the pre-intervention fit, as expected. As in the ArCo method, the PCR approach yields a smaller fraction of cities with significant effects of the price changes. Finally, one important point to highlight is that all three methods suggests that on average the current prices must be decreased.

### 3.3 Effects of Trends

Tables 5.5 and S.6 report the results of the FarmTreat methodology is used without detrending the data in the first step. Compared to the baseline results presented in Tables 6 and 7 in the main text we highlight the following facts. First, the counterfactual model adjustment is similar with only marginal differences concerning the pre-intervention R-squared. Second, without detrending, the average treatment effects are smaller but the rejection rates are higher. Third, the number of municipalities where the estimated $\Delta$ has the correct sign and is statistically significant at the $10 \%$ level is much smaller when we do not include a linear trend in the first step of the methodology, specially in the case of Product V. We note that for this last product, the recommendation is a price increase and not decrease. For the other four products, the
conclusions are similar as the baseline case.

### 3.4 State-Level Aggregation

Tables S. 7 and S.8 report the results of the FarmTreat methodology applied to data aggregated at the state level. The control and treatment groups at the state-level are constructed by aggregating the untreated and treated municipalities in each state, respectively. From the tables we see that for Products IV and V we do not find significant effects at the state level. This is mainly due to heterogeneity across municipalities within each state. On the other hand, for Products I, II and III we find significant effects of price changes on sales. On the average, the optimal price for Products III and V are higher than the actual ones, whereas for Product IV the FarmTreat method indicates that on average the prices should be reduced. However, even for this products the effects are significant in only a fraction of states. These results, corroborates the huge municipality heterogeneity.

### 3.5 Before-and-After Estimation

Table S.9 reports estimation the average treatment effect using the before-and-after estimator. In each panel we report, for each product, the minimum, the $5 \%-, 25 \%-, 50 \%-, 75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities.

## 4 Proof of the Main Result

Before proving our main result, we define below the compatibility constant for convenience.

Definition 1. For $a(n \times n)$ matrix $\boldsymbol{M}$, a set $\mathcal{S} \subseteq[n]$ and a scalar $\zeta \geqslant 0$, the compatibility constant is given by

$$
\begin{equation*}
\kappa(\boldsymbol{M}, \mathcal{S}, \zeta):=\inf \left\{\frac{\left\|\boldsymbol{x}^{T} \boldsymbol{M} \boldsymbol{x}\right\|}{\sqrt{|\mathcal{S}|}}\left\|\boldsymbol{x}_{\mathcal{S}}\right\|_{1}: \boldsymbol{x} \in \mathbb{R}^{n}:\left\|\boldsymbol{x}_{\mathcal{S}^{c}}\right\|_{1} \leqslant \xi\left\|\boldsymbol{x}_{\mathcal{S}}\right\|_{1}\right\} . \tag{S.1}
\end{equation*}
$$

Moreover, we say that $(\boldsymbol{M}, \mathcal{S}, \zeta)$ satisfies the compatibility condition if $\kappa(\boldsymbol{M}, \mathcal{S}, \zeta)>0$.

The compatibility constant is related to $\ell_{1}$-eigenvalue of $\boldsymbol{M}$ restricted to a cone in $\mathbb{R}^{n}$.

### 4.1 Proof of Proposition 1

The fact that $\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}=O_{P}\left(\xi\left|\mathcal{S}_{0}\right|\right)$ follows from Theorem 3 in Fan, Masini, and Medeiros (2021). We are left to show the second part. By the triangle inequality, for $t>T_{0}$ :

$$
\begin{aligned}
\left|\widehat{\alpha}_{t}-\alpha_{t}-V_{t}\right| & =\left|\left(\hat{\boldsymbol{\gamma}}_{1}-\boldsymbol{\gamma}_{1}\right)^{\prime} \boldsymbol{W}_{1 t}+\widehat{\boldsymbol{\lambda}}_{1}^{\prime} \hat{\boldsymbol{F}}_{t}-\boldsymbol{\lambda}_{1}^{\prime} \boldsymbol{F}_{t}+\widehat{\boldsymbol{\theta}}_{1}^{\prime} \hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{\theta}_{1}^{\prime} \boldsymbol{U}_{-1 t}\right| \\
& \leqslant\left|\left(\hat{\boldsymbol{\gamma}}_{1}-\boldsymbol{\gamma}_{1}\right)^{\prime} \boldsymbol{W}_{1 t}\right|+\left|\hat{U}_{1 t}-U_{1 t}\right|+\left|\widehat{\boldsymbol{\theta}}_{1}^{\prime} \hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{\theta}_{1}^{\prime} \boldsymbol{U}_{-1 t}\right| .
\end{aligned}
$$

Using Hölder's inequality, the third term can be further bounded as

$$
\begin{aligned}
\left|\hat{\boldsymbol{\theta}}_{1}^{\prime} \hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{\theta}_{1}^{\prime} \boldsymbol{U}_{-1 t}\right| & \leqslant\left|\hat{\boldsymbol{\theta}}_{1}^{\prime}\left(\hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{U}_{-1 t}\right)\right|+\left|\left(\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right)^{\prime} \boldsymbol{U}_{-1 t}\right| \\
& \leqslant\left\|\hat{\boldsymbol{\theta}}_{1}\right\|_{1}\left\|\hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{U}_{-1 t}\right\|_{\infty}+\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}\left\|\boldsymbol{U}_{-1 t}\right\|_{\infty} \\
& \leqslant\left(\left\|\boldsymbol{\theta}_{1}\right\|_{1}+\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}\right)\left\|\hat{\boldsymbol{U}}_{-1 t}-\boldsymbol{U}_{-1 t}\right\|_{\infty}+\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}\left\|\boldsymbol{U}_{-1 t}\right\|_{\infty} \\
& =O_{P}\left[\left(\left\|\boldsymbol{\theta}_{1}\right\|_{1}+v\left|\mathcal{S}_{0}\right| \psi^{-1}(T)\right) v+v\left|\mathcal{S}_{0}\right| \psi^{-1}(T) \psi^{-1}(n)\right] .
\end{aligned}
$$

Combining the last two expressions we are left with

$$
\left|\widehat{\alpha}_{t}-\alpha_{t}-V_{t}\right| \leqslant\left|\left(\hat{\boldsymbol{\gamma}}_{1}-\boldsymbol{\gamma}_{1}\right)^{\prime} \boldsymbol{W}_{1 t}\right|+\left(1+\left\|\boldsymbol{\theta}_{1}\right\|_{1}+\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}\right)\left\|\hat{\boldsymbol{U}}_{t}-\boldsymbol{U}_{t}\right\|_{\infty}+\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}\left\|\boldsymbol{U}_{t}\right\|_{\infty} .
$$

The first term is $O_{P}(1 / \sqrt{T})$ by Assumption 3 (a). The second is $O_{P}\left(\left|\mathcal{S}_{0}\right| \eta\right)$ because by Assumption 3 (d) we have that $\left\|\boldsymbol{\theta}_{1}\right\|_{1} \leqslant\left|\mathcal{S}_{0}\right|\left\|\boldsymbol{\theta}_{1}\right\|_{\infty} \leqslant C\left|\mathcal{S}_{0}\right|$ and $\left\|\hat{\boldsymbol{\theta}}_{1}-\boldsymbol{\theta}_{1}\right\|_{1}=O_{P}(1)$ under the assumptions of the Proposition. Finally, the third term is $O_{P}\left(\xi\left|\mathcal{S}_{0}\right| n^{1 / p}\right)$ by Assumption 3(b) and the maximum inequality. Therefore we conclude that

$$
\widehat{\alpha}_{t}-\alpha_{t}-V_{t}=O_{P}\left(T^{-1 / 2}+\left|\mathcal{S}_{0}\right| \eta+\xi\left|\mathcal{S}_{0}\right| n^{1 / p}\right)=O_{P}\left[\left|\mathcal{S}_{0}\right|\left(\eta+\xi n^{1 / p}\right)\right] .
$$

Table S.1: Results: Estimation and Inference (ArCo).
The table reports estimation results using the ArCo methodology of Carvalho, Masini, and Medeiros (2018). In each panel we report, for each product, the minimum, the $5 \%-25 \%-, 50 \%-, 75 \%-$, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities. In Panel (a) we report the results for the R-squared of the pre-intervention model. Panel (b) displays the results for the average intervention effect over the experiment period ( $\Delta$ ). Panels (c) and (d) depict the results for the $p$-values of the ressampling test for the null hypothesis $\mathcal{H}_{0}: \delta_{t}=0, \forall t \in\left\{T_{0}+1, \ldots, T\right\}$ using respectively the test statistic $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T} \hat{\delta}_{t}^{2}$ or $\phi\left(\hat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T}\left|\widehat{\delta}_{t}\right|$. Finally, Panel (e) reports the results for the $p$-values for the test for the idiosyncratic contribution.

| Panel (a): R-squared |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| 1 | 0 | 0.1421 | 0.3367 | 0.4389 | 0.6276 | 0.7821 | 0.8958 | 0.4641 | 0.1981 |
| II | 0.4448 | 0.6555 | 0.8691 | 0.9218 | 0.9575 | 0.9851 | 0.9958 | 0.8899 | 0.1073 |
| III | 0.0639 | 0.3119 | 0.4957 | 0.6937 | 0.8181 | 0.9115 | 0.9679 | 0.6554 | 0.2018 |
| IV | 0.3688 | 0.6902 | 0.8823 | 0.9262 | 0.9635 | 0.9888 | 0.9987 | 0.8984 | 0.1056 |
| V | 0 | 0 | 0 | 0.0966 | 0.2210 | 0.4319 | 0.6975 | 0.1452 | 0.1545 |
| Panel (b): Average Treatment Effect (over time): $\Delta$ |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | -20.1194 | -12.0679 | -6.0420 | -2.9966 | -0.6335 | 1.7254 | 7.2911 | -3.6588 | 4.3075 |
| II | -40.6070 | -25.4886 | -9.9769 | -3.1266 | 0.2057 | 9.9614 | 59.7638 | -4.2132 | 11.6643 |
| III | -37.8542 | -8.5142 | -3.3295 | -1.0079 | 0.2364 | 3.7909 | 9.6714 | -2.0799 | 5.8070 |
| IV | -2.5440 | -1.6212 | -0.5723 | 0.1673 | 1.4634 | 3.8332 | 6.4165 | 0.4945 | 1.6339 |
| V | -1.2218 | -0.8548 | -0.4922 | -0.2797 | 0.0044 | 0.4945 | 1.1978 | -0.2476 | 0.4234 |
| Panel (c): $p$-value of the test on squared values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0.0085 | 0.1830 | 0.3702 | 0.6351 | 0.9147 | 0.9787 | 0.4077 | 0.2838 |
| II | 0 | 0.0388 | 0.2273 | 0.4876 | 0.7521 | 0.9521 | 1.0000 | 0.4905 | 0.2967 |
| III | 0 | 0.0306 | 0.2638 | 0.4894 | 0.6638 | 0.8928 | 0.9915 | 0.4735 | 0.2658 |
| IV | 0 | 0 | 0.0888 | 0.3802 | 0.7004 | 0.9029 | 0.9793 | 0.4092 | 0.3162 |
| V | 0 | 0.0894 | 0.3574 | 0.6936 | 0.9149 | 1.0000 | 1.0000 | 0.6452 | 0.3015 |
| Panel (d): $p$-value of the test on absolute values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0.0787 | 0.3149 | 0.5691 | 0.8960 | 0.9872 | 0.3593 | 0.2995 |
| II | 0 | 0.0223 | 0.1818 | 0.5021 | 0.7273 | 0.9504 | 1.0000 | 0.4753 | 0.3095 |
| III | 0 | 0 | 0.2681 | 0.4532 | 0.6766 | 0.8655 | 1.0000 | 0.4624 | 0.2671 |
| IV | 0 | 0 | 0.1033 | 0.3946 | 0.7066 | 0.9318 | 0.9876 | 0.4124 | 0.3220 |
| V | 0 | 0.1234 | 0.3787 | 0.6745 | 0.8809 | 0.9957 | 1.0000 | 0.6284 | 0.2886 |

## Table S.2: Results: Elasticities and Optimal Prices (ArCo).

The table reports elasticities estimates as well the percentage difference between the current prices and the optimal price maximizing profit when the ArCo methodology by Carvalho, Masini, and Medeiros (2018) is used. In each panel we report, for each product, the minimum, the $5 \%-25 \%$-, $50 \%$-, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a given statistic. We consider the distribution over the selected treated municipalities. We only report results concerning the cities where the estimated $\Delta$ has the correct sign and the effects are statistical significance at the $\mathbf{1 0 \%}$ level. The last column indicates the fraction of cities that satisfy the criterium described above. In Panel (a) we report the results for the estimated elasticities. In Panel (b) we show the results for the difference between the current price and the optimal price.

| Panel (a): Elasticities |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| I | -6.1256 | -6.0847 | -3.4902 | -2.7145 | -2.1592 | -1.3585 | -1.2700 | -3.1159 | 1.4785 | 0.1443 |
| II | -16.8229 | -16.8229 | -12.6336 | -11.4970 | -7.5925 | -4.0746 | -4.0746 | -10.5119 | 3.9061 | 0.0882 |
| III | -3.0759 | -3.0759 | -2.8387 | -2.0602 | -1.8896 | -1.6480 | -1.6480 | -2.2876 | 0.5477 | 0.0755 |
| IV | -44.2416 | -34.5020 | -11.9050 | -6.5419 | -4.5606 | -2.4450 | -1.9634 | -10.6109 | 10.1396 | 0.2400 |
| V | -135.5289 | -135.5289 | -19.1707 | -10.1509 | -5.4511 | -4.3575 | -4.3575 | -30.8017 | 51.5716 | 0.0545 |
| Panel (b): Price Discrepancies (\% Difference) |  |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| I | -15.5005 | -15.4441 | -9.3372 | -5.2233 | -0.5061 | 13.6720 | 15.7075 | -4.2981 | 8.5470 | 0.1443 |
| II | -21.3235 | -21.3235 | -20.3202 | -19.9466 | -17.6746 | -12.0246 | -12.0246 | -18.6809 | 2.8620 | 0.0882 |
| III | -9.0999 | -9.0999 | -7.7113 | -1.0822 | 1.1540 | 4.9837 | 4.9837 | -2.4244 | 5.1882 | 0.0755 |
| IV | -18.5830 | -18.2201 | -15.5075 | -12.0222 | -8.7234 | 1.1269 | 5.7524 | -11.2130 | 5.9686 | 0.2400 |
| V | -19.3704 | -19.3704 | -17.1312 | -14.8087 | -10.5669 | -8.2649 | -8.2649 | -14.1585 | 4.1117 | 0.0545 |

## Table S.3: Results: Estimation and Inference (PCR).

The table reports estimation results using principal component regressions. In each panel we report, for each product, the minimum, the $5 \%-, 25 \%-, 50 \%-, 75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities. In Panel (a) we report the results for the R-squared of the pre-intervention model. Panel (b) displays the results for the average intervention effect over the experiment period $(\Delta)$. Panels (c) and (d) depict the results for the $p$-values of the ressampling test for the null hypothesis $\mathcal{H}_{0}: \delta_{t}=0, \forall t \in\left\{T_{0}+1, \ldots, T\right\}$ using respectively the test statistic $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=$ $\sum_{t=T_{0}+1}^{T} \widehat{\delta}_{t}^{2}$ or $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T}\left|\widehat{\delta}_{t}\right|$.

|  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | $5 \%$-quantile | $25 \%$-quantile | Median | $75 \%$-quantile | $95 \%$ quantile | Max | Average | Std. Dev |
| I | 0.1115 | 0.1727 | 0.3307 | 0.4491 | 0.6011 | 0.7294 | 0.7892 | 0.4517 | 0.1707 |
| II | 0.2549 | 0.4633 | 0.7428 | 0.8345 | 0.8815 | 0.9456 | 0.9759 | 0.7898 | 0.1445 |
| III | 0.1026 | 0.1588 | 0.2489 | 0.3466 | 0.5095 | 0.6296 | 0.6944 | 0.3751 | 0.1545 |
| IV | 0.1300 | 0.2384 | 0.5805 | 0.7173 | 0.8236 | 0.8941 | 0.9627 | 0.6723 | 0.1996 |
| V | 0.0255 | 0.0366 | 0.0739 | 0.1236 | 0.2068 | 0.3815 | 0.5079 | 0.1545 | 0.1033 |


| Panel (b): Average Treatment Effect (over time): $\Delta$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | -21.9722 | -17.1898 | -7.6521 | -3.4870 | -1.0735 | 1.6398 | 3.6122 | -5.0798 | 5.6688 |
| II | -47.0186 | -32.5355 | -15.2901 | -7.5150 | -2.8772 | 9.9514 | 40.2040 | -9.2082 | 12.9511 |
| III | -55.4751 | -17.1204 | -7.2165 | -3.4482 | -0.6900 | 1.8316 | 8.8381 | -5.6288 | 9.8650 |
| IV | -4.3269 | -1.9948 | -0.7039 | 0.2394 | 1.5064 | 4.1167 | 7.3901 | 0.5691 | 1.9752 |
| V | -2.0826 | -0.9796 | -0.5058 | -0.1766 | 0.1292 | 0.6744 | 1.6705 | -0.1831 | 0.5190 |


| Product | Panel (c): $p$-value of the test on squared values |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Min | $5 \%$-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| 1 | 0 | 0 | 0.0723 | 0.2553 | 0.6170 | 0.8985 | 0.9915 | 0.3445 | 0.3063 |
| II | 0.0289 | 0.0421 | 0.2355 | 0.4566 | 0.6901 | 0.8983 | 0.9752 | 0.4697 | 0.2844 |
| III | 0 | 0.0664 | 0.2809 | 0.4511 | 0.6340 | 0.9336 | 1.0000 | 0.4624 | 0.2478 |
| IV | 0 | 0.0723 | 0.2459 | 0.4153 | 0.7169 | 0.9917 | 1.0000 | 0.4794 | 0.2878 |
| V | 0 | 0.0596 | 0.3277 | 0.6511 | 0.9319 | 1.0000 | 1.0000 | 0.6050 | 0.3302 |

Panel (d): $p$-value of the test on absolute values

| Product | Min | $5 \%$-quantile | $25 \%$-quantile | Median | $75 \%$-quantile | $95 \%$ quantile | Max | Average | Std. Dev |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I | 0 | 0 | 0.0511 | 0.2128 | 0.6053 | 0.9019 | 0.9957 | 0.3199 | 0.3199 |
| II | 0 | 0.0207 | 0.1570 | 0.4256 | 0.6942 | 0.9298 | 0.9628 | 0.4481 | 0.3038 |
| III | 0 | 0.0102 | 0.2128 | 0.4128 | 0.5957 | 0.9319 | 1.0000 | 0.4248 | 0.2689 |
| IV | 0 | 0.0517 | 0.2149 | 0.4070 | 0.7521 | 0.9690 | 1.0000 | 0.4710 | 0.3001 |
| V | 0 | 0.0511 | 0.2681 | 0.6638 | 0.9234 | 1.0000 | 1.0000 | 0.6084 | 0.3362 |

## Table S.4: Results: Elasticities and Optimal Prices (PCR).

The table reports elasticities estimates as well the percentage difference between the current prices and the optimal price maximizing profit when the counterfactuals are estimated by principal component regression. In each panel we report, for each product, the minimum, the $5 \%-, 25 \%-, 50 \%$-, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a given statistic. We consider the distribution over the selected treated municipalities. We only report results concerning the cities where the estimated $\Delta$ has the correct sign and the effects are statistical significance at the $\mathbf{1 0 \%}$ level. The last column indicates the fraction of cities that satisfy the criterium described above. In Panel (a) we report the results for the estimated elasticities. In Panel (b) we show the results for the difference between the current price and the optimal price.

| Panel (a): Elasticities |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| I | -6.5287 | -5.5838 | -4.3723 | -3.5023 | -2.9266 | -1.7697 | -0.9696 | -3.5689 | 1.1443 | 0.2784 |
| II | -17.7671 | -17.5199 | -14.5000 | -13.1484 | -8.6987 | -2.8126 | -1.9998 | -11.8098 | 4.2565 | 0.1275 |
| III | -3.3805 | -3.3805 | -3.2669 | -2.9047 | -2.7249 | -2.3503 | -2.3503 | -2.9405 | 0.3565 | 0.0755 |
| IV | -15.8735 | -15.8735 | -12.4477 | -11.0990 | -9.2416 | -1.0297 | -1.0297 | -10.3376 | 4.6432 | 0.0700 |
| V | -36.2752 | -36.2752 | -25.0377 | -15.3515 | -5.8105 | -3.3284 | -3.3284 | -16.8591 | 12.4347 | 0.0545 |
| Panel (b): Price Discrepancies (\% Difference) |  |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| I | -16.0044 | -14.6686 | -12.1884 | -9.3865 | -6.5784 | 6.3136 | 27.9067 | -7.6159 | 8.1823 | 0.2784 |
| II | -21.4814 | -21.4382 | -20.8472 | -20.4928 | -18.5425 | -2.0321 | 0.7073 | -18.4545 | 5.8816 | 0.1275 |
| III | -10.5643 | -10.5643 | -10.0456 | -8.1187 | -6.9926 | -4.0815 | -4.0815 | -8.1200 | 2.1928 | 0.0755 |
| IV | -16.5633 | -16.5633 | -15.6963 | -15.2083 | -14.2557 | 28.8451 | 28.8451 | -9.0521 | 16.7294 | 0.0700 |
| V | -18.3610 | -18.3610 | -17.7424 | -16.3414 | -11.1343 | -4.7169 | -4.7169 | -14.1062 | 5.2776 | 0.0545 |

## Table S.5: Results: Estimation and Inference (no trend).

The table reports estimation results without the trend component in the counterfactual model. In each panel we report, for each product, the minimum, the $5 \%$-, $25 \%$-, $50 \%$-, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities aggregated at the state level. In Panel (a) we report the results for the R-squared of the pre-intervention model. Panel (b) displays the results for the average intervention effect over the experiment period ( $\Delta$ ). Panels (c) and (d) depict the results for the $p$-values of the ressampling test for the null hypothesis $\mathcal{H}_{0}: \delta_{t}=0, \forall t \in\left\{T_{0}+1, \ldots, T\right\}$ using respectively the test statistic $\phi\left(\hat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T} \widehat{\delta}_{t}^{2}$ or $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T}\left|\widehat{\delta}_{t}\right|$. Finally, Panel (e) reports the results for the $p$-values for the test for the idiosyncratic contribution.

| Panel (a): R-squared |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0.1112 | 0.1983 | 0.3463 | 0.4910 | 0.6302 | 0.7556 | 0.9029 | 0.4869 | 0.1777 |
| II | 0.4876 | 0.6913 | 0.8721 | 0.9280 | 0.9563 | 0.9850 | 0.9945 | 0.9007 | 0.0905 |
| III | 0.1141 | 0.2904 | 0.5243 | 0.7085 | 0.8324 | 0.9336 | 0.9600 | 0.6736 | 0.2041 |
| IV | 0.3824 | 0.6693 | 0.8802 | 0.9344 | 0.9632 | 0.9869 | 0.9986 | 0.8969 | 0.1101 |
| V | 0.0243 | 0.0378 | 0.0895 | 0.1461 | 0.2706 | 0.4143 | 0.6396 | 0.1876 | 0.1321 |
| Panel (b): Average Treatment Effect (over time): $\Delta$ |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | -16.1305 | -11.3082 | -5.0625 | -2.6195 | -0.8542 | 1.4314 | 9.9071 | -3.3187 | 4.1600 |
| II | -46.3695 | -27.3151 | -10.4665 | -4.1799 | -0.6947 | 6.8649 | 58.5092 | -5.9179 | 11.9431 |
| III | -26.5438 | -9.0437 | -3.0657 | -0.9038 | 0.6108 | 4.8286 | 16.0986 | -1.5804 | 5.2233 |
| IV | -3.9357 | -1.6404 | -0.5186 | 0.2410 | 1.2506 | 4.0381 | 6.3938 | 0.5208 | 1.7143 |
| V | -1.0360 | -0.5738 | -0.2827 | -0.1076 | 0.1859 | 0.7279 | 1.0478 | -0.0468 | 0.3770 |
| Panel (c): $p$-value of the test on squared values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0.0143 | 0.1564 | 0.4213 | 0.6298 | 0.8764 | 0.9872 | 0.4112 | 0.2786 |
| II | 0 | 0.0198 | 0.1818 | 0.4628 | 0.7273 | 0.9793 | 1.0000 | 0.4626 | 0.3138 |
| III | 0 | 0.0170 | 0.2638 | 0.4745 | 0.7064 | 0.9583 | 1.0000 | 0.4864 | 0.2839 |
| IV | 0 | 0 | 0.1302 | 0.3802 | 0.7025 | 0.9545 | 0.9876 | 0.4074 | 0.3110 |
| V | 0 | 0.0766 | 0.3447 | 0.8170 | 0.9872 | 1.0000 | 1.0000 | 0.6779 | 0.3238 |
| Panel (d): $p$-value of the test on absolute values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0.1000 | 0.4000 | 0.5936 | 0.8979 | 0.9957 | 0.3785 | 0.2885 |
| II | 0 | 0.0025 | 0.1446 | 0.4029 | 0.7355 | 0.9694 | 1.0000 | 0.4471 | 0.3213 |
| III | 0 | 0 | 0.2170 | 0.4787 | 0.7234 | 0.9149 | 1.0000 | 0.4757 | 0.2918 |
| IV | 0 | 0 | 0.0992 | 0.3616 | 0.7066 | 0.9360 | 0.9917 | 0.4000 | 0.3137 |
| V | 0 | 0.1064 | 0.4340 | 0.8021 | 0.9915 | 1.0000 | 1.0000 | 0.6974 | 0.3104 |
| Panel (e): p-value of the test for idiosyncratic contribution |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | $25 \%$-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0.0300 | 0.0820 | 0.2525 | 0.6636 | 0.9500 | 0.1824 | 0.2145 |
| II | 0.0080 | 0.0224 | 0.0620 | 0.1250 | 0.2520 | 0.4940 | 0.6760 | 0.1771 | 0.1488 |
| III | 0 | 0 | 0.0080 | 0.0590 | 0.1540 | 0.3468 | 0.5460 | 0.0995 | 0.1180 |
| IV | 0.0300 | 0.0470 | 0.0990 | 0.1920 | 0.2800 | 0.4410 | 0.6400 | 0.2093 | 0.1317 |
| V | 0 | 0.0240 | 0.1160 | 0.2890 | 0.4180 | 0.7080 | 0.8400 | 0.2977 | 0.2083 |

## Table S.6: Results: Elasticities and Optimal Prices (no trend).

The table reports elasticities estimates as well the percentage difference between the current prices and the optimal price maximizing profit. In each panel we report, for each product, the minimum, the $5 \%$-, $25 \%$-, $50 \%$ -, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a given statistic. We consider the distribution over the selected treated municipalities. We only report results concerning the cities where the estimated $\Delta$ has the correct sign and the effects are statistical significance at the $\mathbf{1 0 \%}$ level. The last column indicates the fraction of cities that satisfy the criterium described above. In Panel (a) we report the results for the estimated elasticities. In Panel (b) we show the results for the difference between the current price and the optimal price.

| Panel (a): Elasticities |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| 1 | -6.1709 | -6.1363 | -4.3408 | -2.9859 | -2.2372 | -1.5650 | -1.4268 | -3.4141 | 1.4124 | 0.1753 |
| II | -17.2147 | -16.8507 | -12.9427 | -11.9334 | -8.7978 | -3.4640 | -2.8945 | -10.9383 | 3.8642 | 0.1569 |
| III | -2.8147 | -2.8147 | -2.4532 | -1.8840 | -1.6626 | -1.6254 | -1.6254 | -2.0550 | 0.4905 | 0.0755 |
| IV | -32.7958 | -24.6827 | -11.4079 | -6.6815 | -3.9395 | -2.6159 | -2.4158 | -8.5285 | 7.1821 | 0.2000 |
| V | -30.6356 | -30.6356 | -28.5022 | -25.0506 | -23.6188 | -15.6706 | -15.6706 | -24.7547 | 5.2214 | 0.0545 |
| Panel (b): Price Discrepancies (\% Difference) |  |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| I | -15.5604 | -15.5143 | -12.1436 | -6.9179 | -1.3130 | 8.7227 | 11.3806 | -6.3656 | 7.4661 | 0.1753 |
| II | -21.3911 | -21.3250 | -20.4324 | -20.1047 | -18.5991 | -9.0741 | -7.0214 | -18.6833 | 3.6201 | 0.1569 |
| III | -7.5915 | -7.5915 | -4.5816 | 1.2230 | 4.7269 | 5.4073 | 5.4073 | 0.0690 | 5.2832 | 0.0755 |
| IV | -18.1886 | -17.4421 | -15.3303 | -12.2240 | -7.0189 | -0.4867 | 0.9835 | -10.4346 | 5.6478 | 0.2000 |
| V | -18.1073 | -18.1073 | -17.9851 | -17.7375 | -17.6224 | -16.5487 | -16.5487 | -17.6231 | 0.5603 | 0.0545 |

## Table S.7: Results: Estimation and Inference (state level).

The table reports estimation results at the state level. In each panel we report, for each product, the minimum, the $5 \%-25 \%-, 50 \%$-, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities aggregated at the state level. In Panel (a) we report the results for the R-squared of the pre-intervention model. Panel (b) displays the results for the average intervention effect over the experiment period ( $\Delta$ ). Panels (c) and (d) depict the results for the $p$-values of the ressampling test for the null hypothesis $\mathcal{H}_{0}: \delta_{t}=0, \forall t \in\left\{T_{0}+1, \ldots, T\right\}$ using respectively the test statistic $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T} \widehat{\delta}_{t}^{2}$ or $\phi\left(\widehat{\delta}_{T_{0}+1}, \ldots, \widehat{\delta}_{T}\right)=\sum_{t=T_{0}+1}^{T}\left|\widehat{\delta}_{t}\right|$. Finally, Panel (e) reports the results for the $p$-values for the test for the idiosyncratic contribution.

| Panel (a): R-squared |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0.3553 | 0.3804 | 0.6776 | 0.8027 | 0.8969 | 0.9573 | 0.9603 | 0.7593 | 0.1814 |
| II | 0.8830 | 0.8895 | 0.9410 | 0.9812 | 0.9934 | 0.9962 | 0.9962 | 0.9659 | 0.0351 |
| III | 0.2983 | 0.3604 | 0.6552 | 0.7726 | 0.8651 | 0.9422 | 0.9642 | 0.7329 | 0.1763 |
| IV | 0.7566 | 0.8014 | 0.9377 | 0.9684 | 0.9874 | 0.9952 | 0.9962 | 0.9480 | 0.0587 |
| V | 0.0996 | 0.1249 | 0.1795 | 0.3048 | 0.5028 | 0.8898 | 0.9024 | 0.3687 | 0.2298 |
| Panel (b): Average Treatment Effect (over time): $\Delta$ |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | -4.9172 | -4.6035 | -3.3475 | -1.8782 | 0.1261 | 1.8318 | 1.8886 | -1.6650 | 2.1306 |
| II | -17.5812 | -16.8032 | -9.1995 | -2.6898 | 2.0167 | 13.5124 | 14.5276 | -2.9574 | 9.0027 |
| III | -7.5112 | -6.7144 | -3.5990 | -0.9333 | 0.3148 | 14.7754 | 32.3728 | -0.2965 | 7.6855 |
| IV | -2.0756 | -1.7365 | -0.7512 | -0.3154 | 0.3816 | 0.8821 | 1.0061 | -0.2839 | 0.7757 |
| V | -0.7695 | -0.6821 | -0.3421 | -0.1571 | -0.0078 | 0.3294 | 0.4216 | -0.1837 | 0.2904 |
| Panel (c): $p$-value of the test on squared values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0.0511 | 0.2511 | 0.4883 | 0.8736 | 0.9957 | 0.3041 | 0.2823 |
| II | 0 | 0 | 0.0548 | 0.3017 | 0.5610 | 0.9731 | 0.9876 | 0.3717 | 0.3427 |
| III | 0 | 0.0194 | 0.2043 | 0.4128 | 0.7511 | 0.9387 | 0.9830 | 0.4503 | 0.3201 |
| IV | 0.0331 | 0.0793 | 0.2500 | 0.4215 | 0.5723 | 0.7901 | 0.8017 | 0.4166 | 0.2206 |
| V | 0 | 0.0070 | 0.3170 | 0.8426 | 0.9543 | 0.9930 | 1.0000 | 0.6470 | 0.3741 |
| Panel (d): $p$-value of the test on absolute values |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0.0032 | 0.3404 | 0.4553 | 0.9109 | 0.9957 | 0.3189 | 0.2930 |
| II | 0 | 0 | 0.0207 | 0.3430 | 0.4824 | 0.9599 | 0.9628 | 0.3493 | 0.3421 |
| III | 0 | 0 | 0.1372 | 0.3830 | 0.7660 | 0.9257 | 0.9617 | 0.4213 | 0.3406 |
| IV | 0.0331 | 0.0605 | 0.2397 | 0.4793 | 0.5981 | 0.7837 | 0.8140 | 0.4256 | 0.2377 |
| V | 0 | 0 | 0.3319 | 0.7830 | 0.9340 | 0.9930 | 1.0000 | 0.6237 | 0.3796 |
| Panel (e): $p$-value of the test for idiosyncratic contribution |  |  |  |  |  |  |  |  |  |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev |
| I | 0 | 0 | 0 | 0.0020 | 0.0210 | 0.1724 | 0.1920 | 0.0314 | 0.0593 |
| II | 0 | 0 | 0.0110 | 0.0380 | 0.0545 | 0.2964 | 0.3440 | 0.0616 | 0.0893 |
| III | 0 | 0 | 0.0105 | 0.0420 | 0.0990 | 0.5157 | 0.6080 | 0.1017 | 0.1640 |
| IV | 0.0940 | 0.1164 | 0.1780 | 0.2040 | 0.2310 | 0.5657 | 0.6420 | 0.2408 | 0.1292 |
| V | 0 | 0 | 0.0350 | 0.0820 | 0.1350 | 0.5735 | 0.6780 | 0.1370 | 0.1767 |

## Table S.8: Results: Elasticities and Optimal Prices (state level).

The table reports elasticities estimates as well the percentage difference between the current prices and the optimal price maximizing profit. In each panel we report, for each product, the minimum, the $5 \%$-, $25 \%$-, $50 \%$ -, $75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a given statistic. We consider the distribution over the selected treated municipalities. We only report results concerning the cities where the estimated $\Delta$ has the correct sign and the effects are statistical significance at the $\mathbf{1 0 \%}$ level. The last column indicates the fraction of cities that satisfy the criterium described above. In Panel (a) we report the results for the estimated elasticities. In Panel (b) we show the results for the difference between the current price and the optimal price.

| Panel (a): Elasticities |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Product | Min | 5\%-quantile | 25\%-quantile | Median | 75\%-quantile | 95\% quantile | Max | Average | Std. Dev | Fraction |
| 1 | -1.8282 | -1.8282 | -1.7647 | -1.5235 | -1.1270 | -0.8921 | -0.8921 | -1.4431 | 0.3720 | 0.2222 |
| II | -11.5235 | -11.5235 | -7.7493 | -5.8728 | -4.4528 | -4.1237 | -4.1237 | -6.5147 | 2.9592 | 0.1852 |
| III | -3.2089 | -3.2089 | -2.9427 | -1.9333 | -1.0804 | -0.9708 | -0.9708 | -2.0116 | 1.1006 | 0.1481 |
| IV | - | - | - | - | - | - | - | - | - | - |
| V | - | - | - | - | - | - | - | - | - | - |

Panel (b): Price Discrepancies (\% Difference)

| Product | Min | $5 \%$-quantile | $25 \%$-quantile | Median | $75 \%$-quantile | $95 \%$ quantile | Max | Average | Std. Dev | Fraction |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I | 3.6865 | 3.6865 | 4.6703 | 9.3255 | 20.7040 | 32.3850 | 32.3850 | 13.3495 | 11.2129 | 0.2222 |
| II | -19.9566 | -19.9566 | -17.4338 | -15.7817 | -13.0450 | -12.1705 | -12.1705 | -15.5676 | 3.0364 | 0.1852 |
| III | -9.7734 | -9.7734 | -8.2241 | 4.9918 | 21.4037 | 26.1490 | 26.1490 | 6.5898 | 17.5845 | 0.1481 |
| IV | - | - | - | - | - | - | - | - | - | - |
| V | - | - | - | - | - | - | - | - | - | - |

Table S.9: Results: Estimation and Inference (Before-and-After).
The table reports estimation the average treatment effect using the before-and-after estimator. In each panel we report, for each product, the minimum, the $5 \%-, 25 \%-, 50 \%-, 75 \%$-, and $95 \%$-quantiles, maximum, average, and standard deviation for a variety of different statistics. We consider the distribution over the treated municipalities.

| Product | Min | $5 \%$-quantile | $25 \%$-quantile | Median | $75 \%$-quantile | $95 \%$ quantile | Max | Average | Std. Dev |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I | -23.8652 | -17.2270 | -8.1333 | -4.1126 | -1.1093 | 2.1760 | 11.5150 | -5.2622 | 6.0399 |
| II | -74.8229 | -53.2274 | -30.7149 | -18.4681 | -10.3370 | 1.8621 | 13.1138 | -22.0736 | 16.9785 |
| III | -48.8512 | -15.3860 | -5.6494 | -2.1679 | -0.5397 | 2.4336 | 11.1025 | -3.9888 | 7.0311 |
| IV | -5.5069 | -4.7638 | -2.1703 | -1.2016 | -0.1093 | 1.7353 | 3.5901 | -1.2483 | 1.9274 |
| V | -2.0595 | -1.3942 | -0.8139 | -0.4505 | -0.1244 | 0.4032 | 1.1809 | -0.4682 | 0.5426 |

## Figure S.1: Data for Product II.

Panel (a) reports the daily sales divided by the number of stores aggregated for all cities as well as for the treatment and control groups. The plot also indicates the date of the intervention. Panels (b) and (c) display the distribution of the average sales per store over time across municipalities in the treatment and control groups, respectively. Panels (d) and (e) present fan plots of sales across municipalities in the treatment and control groups for each given time point. The black curves represent the cross-sectional mean over time and the vertical green line indicates the date of intervention.

Product II






Figure S.2: Data for Product III.
Panel (a) reports the daily sales divided by the number of stores aggregated for all cities as well as for the treatment and control groups. The plot also indicates the date of the intervention. Panels (b) and (c) display the distribution of the average sales per store over time across municipalities in the treatment and control groups, respectively. Panels (d) and (e) present fan plots of sales across municipalities in the treatment and control groups for each given time point. The black curves represent the cross-sectional mean over time and the vertical green line indicates the date of intervention.


## Figure S.3: Data for Product IV.

Panel (a) reports the daily sales divided by the number of stores aggregated for all cities as well as for the treatment and control groups. The plot also indicates the date of the intervention. Panels (b) and (c) display the distribution of the average sales per store over time across municipalities in the treatment and control groups, respectively. Panels (d) and (e) present fan plots of sales across municipalities in the treatment and control groups for each given time point. The black curves represent the cross-sectional mean over time and the vertical green line indicates the date of intervention.


Figure S.4: Data for Product V.
Panel (a) reports the daily sales divided by the number of stores aggregated for all cities as well as for the treatment and control groups. The plot also indicates the date of the intervention. Panels (b) and (c) display the distribution of the average sales per store over time across municipalities in the treatment and control groups, respectively. Panels (d) and (e) present fan plots of sales across municipalities in the treatment and control groups for each given time point. The black curves represent the cross-sectional mean over time and the vertical green line indicates the date of intervention.


## Figure S.5: Results for Product II

Panel (a) displays a fan plot, across $n_{1}$ municipalities in the treatment group, of the $p$-values of the re-sampling test for the null $\mathscr{H}_{0}: \delta_{t}=0$ at each time $t$ after the treatment. The black curve represents the median $p$-value across municipalities over $t$. Panel (b) shows an example for one municipality. The panel depicts the actual and counterfactual sales per store for the post-treatment period. $95 \%$ confidence intervals for the counterfactual path is also displayed.


Figure S.6: Results for Product III

Panel (a) displays a fan plot, across $n_{1}$ municipalities in the treatment group, of the $p$-values of the re-sampling test for the null $\mathscr{H}_{0}: \delta_{t}=0$ at each time $t$ after the treatment. The black curve represents the median $p$-value across municipalities over $t$. Panel (b) shows an example for one municipality. The panel depicts the actual and counterfactual sales per store for the post-treatment period. $95 \%$ confidence intervals for the counterfactual path is also displayed.

Product III


Figure S.7: Results for Product IV
Panel (a) displays a fan plot, across $n_{1}$ municipalities in the treatment group, of the $p$-values of the re-sampling test for the null $\mathscr{H}_{0}: \delta_{t}=0$ at each time $t$ after the treatment. The black curve represents the median $p$-value across municipalities over $t$. Panel (b) shows an example for one municipality. The panel depicts the actual and counterfactual sales per store for the post-treatment period. $95 \%$ confidence intervals for the counterfactual path is also displayed.


Figure S.8: Results for Product V
Panel (a) displays a fan plot, across $n_{1}$ municipalities in the treatment group, of the $p$-values of the re-sampling test for the null $\mathscr{H}_{0}: \delta_{t}=0$ at each time $t$ after the treatment. The black curve represents the median $p$-value across municipalities over $t$. Panel (b) shows an example for one municipality. The panel depicts the actual and counterfactual sales per store for the post-treatment period. $95 \%$ confidence intervals for the counterfactual path is also displayed.

Product V



Figure S.9: Daily Inventory Distribution.


## References

Carvalho, C., R. Masini, and M. Medeiros (2018): "ArCo: An Artificial Counterfactual Approach for High-Dimensional Panel Time-Series Data," Journal of Econometrics, 207, 352-380.

Fan, J., R. Masini, and M. Medeiros (2021): "Bridging Factor and Sparse Models," arxiv:2102.11341, Princeton University.

Gobillon, L., and T. Magnac (2016): "Regional Policy Evaluation: Interactive Fixed Effects and Synthetic Controls," Review of Economics and Statistics, 98, 535-551.

